

An Integrated Model of Academic Self-Concept Development: Academic Self-Concept, Grades, Test Scores, and Tracking Over 6 Years

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Our newly proposed integrated academic self-concept model integrates 3 major theories of academic self-concept formation and developmental perspectives into a unified conceptual and methodological framework. Relations among math self-concept (MSC), school grades, test scores, and school-level contextual effects over 6 years, from the end of primary school through the first 5 years of secondary school (a representative sample of 3,370 German students, 42 secondary schools, 50% male, *M* age at grade 5 = 11.75) support the (1) internal/external frame of reference model: Math school grades had positive effects on MSC, but the effects of German grades were negative; (2) reciprocal effects (longitudinal panel) model: MSC was predictive of and predicted by math test scores and school grades; (3) big-fish-little-pond effect: The effects on MSC were negative for school-average achievement based on 4 indicators (primary school grades in math and German, school-track prior to the start of secondary school, math test scores in the first year of secondary school). Results for all 3 theoretical models were consistent across the 5 secondary school years: This supports the prediction of developmental equilibrium. This integration highlights the robustness of support over the potentially volatile early to middle adolescent period; the interconnectedness and complementarity of 3 ASC models; their counterbalancing strengths and weaknesses; and new theoretical, developmental, and substantive implications at their intersections.

Keywords: developmental equilibrium, math self-concept, frame of reference effects, reciprocal effects, big-fish-little-pond effects

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Self-concept and related self-beliefs are key constructs in developmental and educational psychology. For many developmental researchers, and in many early childhood programs (e.g., Fantuzzo et al., 1996), self-concept has been a “cornerstone of both social and emotional development” (Kagan, Moore, & Bredekamp, 1995, p. 18; also see Davis-Kean & Sandler, 2001; Marsh, Ellis, & Craven, 2002). Academic self-concept (ASC) is also widely accepted as a critical psychological construct that leads to success in educational settings (Chen, Yeh, Hwang, & Lin, 2013; Marsh & Craven, 2006; Marsh & Yeung, 1997), in social and emotional situations (Harter, 2012;

Marsh, Parada, Craven, & Finger, 2004, Pekrun, 2006), and in daily life more generally (Eccles, 2009; Elliot & Dweck, 2005).

Before outlining the integrated model, we begin with a brief overview of the contributory theories, three of the most important theoretical models in ASC research, which are as follows:

- The internal/external frame of reference (I/E) model relates math and verbal achievement to corresponding measures of ASC;
- the reciprocal effects model (REM) of relations between academic achievement and ASC over time; and

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- the big-fish-little-pond effect (BFLPE), that is the negative effect of school-average achievement on ASC.

Although there is much support for each of these theoretical models considered separately, to our knowledge no study has considered all three models within a unified theoretical framework and a single statistical model incorporating parameter estimates to test all three models simultaneously, using a database suitable for testing all three within a single integrated model.

Historically, the understanding of ASC has been limited by the piecemeal approaches that are endemic when separate theories are considered each in isolation. Thus, for example, the main focus of the I/E model is the juxtaposition of math and verbal constructs; however, this focus on domain specificity is largely ignored in the other two models, which typically are tested within a single academic domain. Likewise, the focus of the reciprocal effects model is on longitudinal relations of ASC and achievement over time, but this longitudinal perspective is largely ignored by the other two models, which typically are tested using cross-sectional data. The critical feature of the BFLPE model is its multilevel consideration of contextual effects (the effects of school-average achievement on self-concept), but this multilevel perspective is largely ignored by the other two models, which typically are tested with single-level models. Importantly, this integration of ASC theories (hereafter referred to as the integrated ASC model) results in a number of new predictions (see the online Supplemental Materials, Section 7, for new predictions that could not be derived from the individual components of the integrated ASC model when considered separately). To this integration of models we add a developmental perspective, in which we evaluate support for the consistency of effects (which we subsequently refer to as *developmental equilibrium*) across the potentially turbulent, a period of early to middle adolescence (first 5 years of secondary school) that involves so many biological and psychological changes (e.g., Eccles, 2009; Eccles et al., 1993; Harter, 2012; Steinberg, 2008).

We also note that the integrated ASC model, and empirical tests of the model, are important because each of the three separate theories leaves open the question as to whether the effects hypothesized in one model are independent of the effects hypothesized in the other two. Only the integrated ASC model (and tests of this model) allows us to examine the robustness of the effects considering all predicted effects combined.

The new integrated ASC model (see Figure 1) incorporates the I/E, REM, and BFLPE models, which are based on the theoretical and empirical work of Marsh and colleagues (e.g., Marsh, 2007; Marsh & Craven, 2006; Marsh, Seaton, et al., 2008; Möller & Marsh, 2013). Thus, its overarching aim is to systematically explain the relations between ASC and academic achievement across domains (dimensional comparisons within the I/E model; i.e., “My accomplishments in one domain relative to accomplishments in other domains”), time (development within the REM model; i.e., “My current accomplishments relative to past accomplishments”), and school peer group (social comparisons within the BFLPE model; i.e., “My accomplishments relative to those of my peer group”).

In summary, there are important advantages in bringing these three theoretical models within the unified framework proposed here. The complementarity of these different theoretical perspectives allows us to achieve a broader understanding of the formation of ASC. In addition, integrating the three into a single unified

framework results in new theoretical predictions arising from the intersections of the different models (see subsequent discussion). Methodologically, it is also important to emphasize that with appropriate data, all three models can be tested within a single statistical model. We demonstrate how parameter estimates based on one unified statistical model provide tests of each of the three ASC models from a developmental perspective: This reinforces their complementarity.

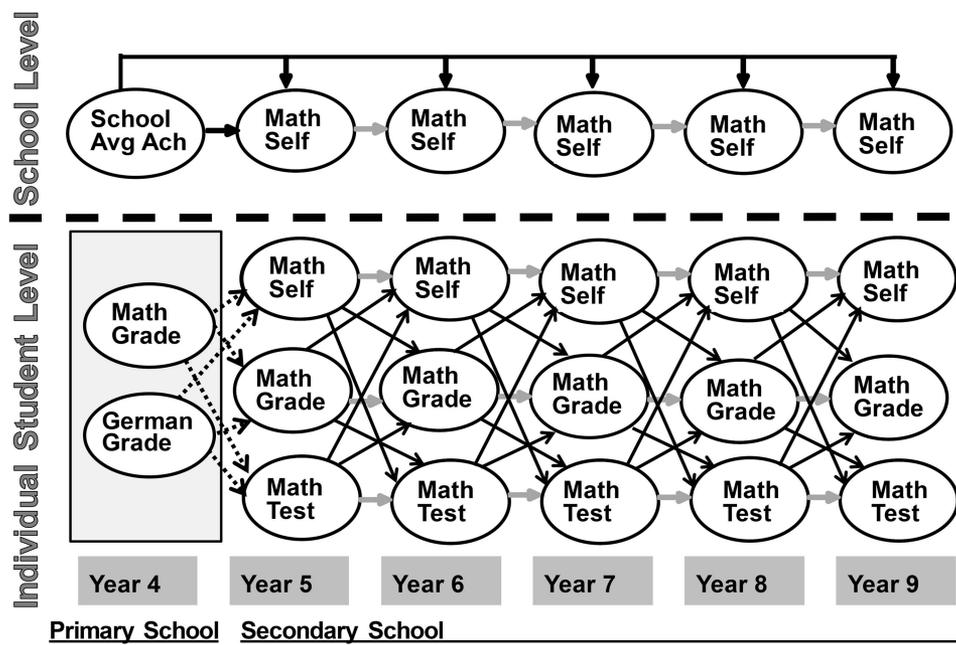
Integration of Three Theoretical Models of ASC Formation

In the present investigation, we aim to investigate how students develop their beliefs about their competence throughout their adolescence. In pursuing this aim, we take into account three main influences that have been identified in prior research in the three theoretical models, which are as follows: dimensional comparison, reciprocal effects, and social comparison effects.

The I/E Model: Dimensional Comparison Effects

ASCs in specific academic domains are much more differentiated than are the corresponding measures of achievement. Indeed, even though math and verbal achievements tend to be highly correlated, math and verbal self-concepts tend to be nearly uncorrelated (Marsh, 1986, 2007; Marsh, Kuyper, Seaton, et al., 2014). The I/E model provides a theoretical rationale for these seemingly paradoxical results, in positing that ASC in a particular school subject is formed in relation to two comparison processes: an external (*social comparison*) reference, in which students compare their performances in a particular school subject with the performances of other students in the same school subject, and an internal (*dimensional comparison*) reference, in which students compare their own performances in that particular school subject with their own performances in other school subjects. In particular, Marsh (1986) proposed that students use an internal comparison process, whereby academic achievement in one domain (e.g., verbal) provides a frame of reference for forming ASC in a contrasting domain (e.g., math).

Although the I/E model posits that achievement is highly positively predictive of ASC in the matching domain, the critical theoretical predictions are the negative cross-paths leading from achievement in one subject to ASC in the other subject; for example, verbal achievement to math self-concept (MSC). The theoretical rationale for the negative cross-paths (*dimensional comparisons*) is that students will use verbal achievement, for example, as a basis of comparison in the formation of their MSC. Thus, high verbal achievement will detract from a high math self-concept; likewise, students who have good math achievement will have lower MSCs if their verbal achievement is much higher than their math achievement. Following initial tests of the I/E model (Marsh, 1986), the I/E model predictions were found to be supported in 26 countries using Programme for International Student Assessment (PISA) data (Marsh & Hau, 2004). In a subsequent meta-analysis based on 69 data sets, Möller et al. (2009) reported that math and verbal achievements were highly correlated ($r = .67$), but that the corresponding self-concepts were nearly uncorrelated ($r = .10$). The paths from math achievement to MSC were positive ($\beta = .61$), but paths from verbal achievement to MSC were negative ($\beta = -.27$).



- Horizontal (stability) paths from one construct to the same construct in an subsequent Year (only)
- Cross-paths from one construct to a different construct (only lag-1 paths shown)
- Tests of the internal/external frame of reference model. Paths from primary school grades to secondary school outcomes (only lag-1 paths shown)

Figure 1. Integrated academic self-concept model. This single conceptual model (and associated statistical model in Table 1) provides tests of all three theoretical models of academic self-concept (ASC) and developmental equilibrium. The paths from math and German grades in Year 4 to MSC in Years 5 through 9 provide tests of the internal/external frame of reference model. The paths relating math achievement (test scores and grades) and self-concept in Years 5 through 8 to these measures in subsequent years provide tests of the reciprocal effects model. At the school level, the paths from school-average achievement to MSC provide tests of the big-fish-little-pond effect. The consistency of the paths over time provides tests of developmental equilibrium. For the purposes of illustration, only lag-1 paths are shown (effects of each variable on variables in the immediately subsequent wave). However, in the final a priori model (see the integrated ASC model, Table 1; also see the online Supplemental Materials, Section 3) the following are shown: (i) paths from both math and German grades in Year 4 to all Years 5 through 9 outcomes (the dashed lines from Year 4 school grades to Year 5, but also paths for Year 4 school grades to outcomes in Years 6 through 9); (ii) all lag-1 to lag-4 autocorrelation (horizontal) test-retest paths relating all Years 5 through 9 variables in each Year to the same variable in all subsequent years (the solid horizontal lines from each variable in Years 5 through 8 to the same variable in the next year, lag-1 paths, but also paths from each variable to the same variable in all subsequent years: lags 2 through 4). These paths are constrained to be invariant across years (e.g., lag-1 paths from test scores in Year 5 to test scores in Year 6 are the same as the lag-1 path from test scores Year 8 to test scores in Year 9); (iii) All lag-1 and lag-2 cross-paths relating all Years 5 through 9 variables in each year to each of the different variables in the next year (only lag-1 paths are shown). These paths were constrained to be invariant across years (e.g., lag-1 paths from test scores in Year 5 to MSC in Year 6 were the same as the lag-1 path from test scores in Year 8 to MSC in Year 9); and (iv) covariances between all variables measured within the same wave (e.g., math and German grades at Year 4; MSC, test scores, and grades at Year 5). Support for developmental equilibrium is based on goodness of fit tests (presented in greater detail in the online Supplemental Materials, Section 3) that provide support for this a priori prediction and a statistical basis for constraining paths to be invariant across years. Parameter estimates (see Table 1) from this one multilevel, longitudinal path model are used to test the internal/external frame of reference model, the reciprocal effects model and the big-fish-little-pond effect, in which school-average achievement is based on various combinations of Year 4 school grades, school track, and Year 5 test scores. The Mplus syntax and output showing the full set of lagged paths are presented in the Supplemental Materials, Section 8.

The REM of Relations Between ASC and Achievement

ASC and academic achievement are substantially correlated, but a critical question with important theoretical and policy–practice implications is the temporal ordering of these constructs. Traditional approaches to this issue (Calsyn & Kenny, 1977) took an “either-or” approach—either prior achievement leads to subsequent ASC (a skill development model) or prior ASC leads to subsequent achievement (a self-enhancement model). However, integrating theoretical and statistical perspectives, Marsh (1990) argued for a dynamic reciprocal effects model (REM) that incorporates both the skill development and the self-enhancement models, such that both ASC and achievement are posited to be causes and also effects of each other; the REM is testable when both constructs are collected in at least two but preferably three or more waves of data.

In meta-analyses of REM studies, Valentine et al. (2004; also see Huang, 2011) found consistent support for the REM. It is not surprising that prior achievement has an effect on ASC, as this is consistent with ASC theory and research. However, Valentine et al. demonstrated that the effect of prior ASC on subsequent achievement, after controlling for the effects of prior achievement, was also highly significant overall, and positive in 90% of the studies they considered. However, two limitations of the results summarized in these meta-analyses are addressed here.

First, although REM studies are necessarily longitudinal, most studies are based on one, two or, perhaps, three waves of data, and do not cover an extended developmental period. Here we evaluate support for the REM on the basis of six waves of data covering the early to middle adolescent (late primary school through high school; see subsequent discussion of Figure 1) period: This provides a stronger test of the consistency of effects over the potentially turbulent developmental period.

Second, in REM studies, achievement typically is assessed by standardized tests or school grades—and yet, the different achievement indicators have different implications. School grades are a particularly salient source of feedback to students and their parents, are easily compared among classmates, and have important implications for academic careers. Hence, school grades tend to be more correlated with ASCs than they are with test scores (e.g., Marsh, Kuyper, Morin, et al., 2014; Marsh, Kuyper, Seaton, et al., 2014; Marsh, Trautwein, et al., 2005). However, school grades typically are idiosyncratic to specific teachers, settings, and schools. In particular, teachers typically grade on a curve, allocating the highest and lowest grades to the relatively best and least-well performing students within a classroom, respectively. Hence, teachers use the classroom as a narrow frame of reference in their grading procedure, largely ignoring the absolute levels of achievement of students in their class based on a common metric that generalizes over all students. Although the classic meta-analyses support REM predictions in respect of both school grades and test scores, most individual studies have included only one of these indicators of achievement, and apparently none have juxtaposed the two over such an extended developmental period as that considered here or in relation to developmental perspectives. Furthermore, incorporating both verbal and math test scores into the integrated ASC model integrates the typically cross-sectional tests of the I/E model with the reciprocal effects inherent in the REM,

and the multilevel effects of school-average achievement in BFLPE studies.

The BFLPE: Social Comparison Effects

According to the BFLPE, students compare their own academic achievement with the achievements of their classmates, and use this social comparison as the basis of their ASCs (Marsh, Kuyper, Morin, et al., 2014; Marsh, Abduljabbar, et al., 2015; Marsh & Parker, 1984; Marsh, Seaton, et al., 2008; Nagengast & Marsh, 2012; Zell & Alicke, 2009). In the BFLPE, students who attend high-ability schools tend to have lower ASCs than do equally able students who attend mixed- or low-ability schools, which is a negative effect of school-average achievement on ASC.

There is now considerable support for the negative effects of school-average achievement on ASC (see reviews by Marsh, Seaton, et al., 2008; Marsh & Seaton, 2015). Demonstrating that the BFLPE is one of psychology’s most cross-culturally universal phenomena, three successive Programme for International Student Assessment (PISA) data collections (Marsh & Hau, 2003: 103,558 students from 26 countries; Seaton, Marsh, & Craven, 2010: 265,180 students from 41 countries; Nagengast & Marsh, 2012: 397,500 students from 57 countries) showed that the effect of school-average achievement on ASC was negative in all but one of the 123 samples, and significantly so in 114 samples. Further, not only does the BFLPE tend to increase in size during the period that students attend the same high school (Marsh, Köller, & Baumert, 2001), but Marsh, Trautwein, Lüdtke, Baumert, and Köller (2007) have shown that the BFLPE formed in high school is equally large, or larger, two (Study 1) and four (Study 2) years after graduation from high school. Similarly, Marsh and O’Mara (2008; Guo, Marsh, Parker, & Morin, 2015) showed that ASC formed in high school contributed to the prediction of long-term educational attainment eight years later, and beyond the effects of school grades, standardized achievement tests, IQ, and socioeconomic status. Nevertheless, BFLPE studies have several limitations; these are addressed in the present research.

First, most BFLPE studies are cross-sectional, so that the school-average achievement associated with a particular school might reflect either the ability of students prior to attending that school, or the subsequent effects of the school on achievement—potentially confounding the temporal ordering. Second, the primary focus on school-average achievement as a measure of de facto selectivity (e.g., selectivity based on neighborhood) might not generalize to explicit selectivity when students are tracked into different schools on the basis of prior achievement (see Marsh, Köller, & Baumert, 2001). Third, because school grades typically do not reflect a common metric, previous BFLPE studies have been based on test scores rather than on school grades. However, school grades are a more salient measure of achievement than are standardized test scores, influence ASCs more than test scores do, and are often an important basis for tracking students. Nevertheless, because of the grading on a curve effect, school-average grades are unlikely to be comparable over schools when schools differ in terms of mean achievement levels. Finally, BFLPE studies traditionally evaluate school-average achievement in the same domain as the corresponding measure of ASC (e.g., the effects of school-average math achievement on MSC). However, integrating the I/E and BFLPE leads to the question of how school-average

achievement in a contrasting domain might affect ASC (i.e., the effect of school-average verbal achievement on MSC). Given the unique design of our study, we were able to address each of these limitations of prior BFLPE research by integrating different theoretical and developmental perspectives.

Integrating Developmental Equilibrium into the Integrated ASC Model

Developmental equilibrium. Our study is based on testing the empirical support for what has been referred to as developmental equilibrium (Marsh, Craven, et al., 2016; Marsh, Pekrun, Lichtenfeld, et al., 2016; Marshall et al., 2015; see the online Supplemental Materials, Section 1, for further discussion). Although the term *equilibrium* comes from the physical sciences, it has long been applied as an analogous concept in psychological theorizing—particularly in developmental psychology. Equilibrium is reached when a system achieves a state of balance between potentially counterbalancing, opposing forces (e.g., physiological homeostasis, Cannon, 1932; psychological equilibrium balancing competing drives and desires, Argyle, 1967; Erikson, 1974; self-actualization as an equilibrium between actual and ideal self-perceptions; Rogers, 1961). In child development, Piaget and Cook (1952) argued that the psychological system aims to achieve a steady state of equilibrium that allows children to accommodate new experiences using existing schemas, whereas disequilibrium forces children to change their cognitive structures to regain equilibrium. In each of these different perspectives on equilibrium, the critical issue is that of balance, posited to be a psychologically desirable state, and indicating consistency over time.

Here we evaluate support for developmental equilibrium through tests of the consistency of relations among critical variables over early to middle adolescence—that is, whether the self-system is consistently in a state of balance during this period. Thus, for example, Davis-Kean et al. (2008; also see Davis-Kean, Jager, & Collins, 2009) reported that the relation between ASC and achievement changed with age for young children, but became relatively stable from the age of about 12. This suggests that this relation is stable and has reached a state of equilibrium during the early to middle adolescent period, which is the focus of our study.

Although the term is often used metaphorically, achieving a state of equilibrium clearly has important substantive and psychological implications. Here, however, we operationalize this perspective by integrating it with formal statistical models of longitudinal invariance, based on models of the invariance of effects across multiple waves of data. This has theoretical, developmental, and substantive implications: for example, the question of whether the effect sizes of critical components in each of these models of ASC formation vary developmentally (Eccles, 2009; Marsh, 2007; Marsh & O'Mara, 2008; Marsh, Seaton, et al., 2008; Murayama et al., 2013), including the relative sizes of paths leading from achievement to ASC, and ASC to achievement in the REM; the size of the BFLPE; and the strength of the internal comparison process in the I/E model. In summary, our major developmental focus is on tests of developmental equilibrium: the consistency over time of relations between ASC and achievement in reference to predictions from the three theoretical ASC models, over the potentially turbulent period of early to middle adolescence (e.g., Eccles, 2009; Eccles et al., 1993; Harter, 2012; Steinberg, 2008).

We also note that our notion of developmental equilibrium closely resembles Fraley, Roisman, and Haltigan's (2013) "Legacy of Early Experiences in Development," which they present as an important, ongoing debate in developmental science. Specifically, they argue: "By studying the pattern of associations across time, it should be possible to gain greater insight into the legacy of early experiences" (p. 113). Indeed, their paradigmatic models closely resemble our integrated ASC model (see Figure 1). They proposed models of the longitudinal effects of a particular event in time that are similar to our evaluation of primary school grades and school-average ability. Their emphasis, like ours, was on the direct and indirect effects of a variable over time. However, as in our evaluation of the REM, they also proposed cross-lagged panel models of the same variables measured on multiple occasions over time. Indeed, our a priori hypothesis of developmental equilibrium can be seen as a special case of a more general model, in which selected effects are consistent over time—a possibility that they introduced by testing the equality of parameter estimates across multiple waves of data. Like us they argued for consideration of more than two waves of data in which the same constructs are studied—ideally, covering an important developmental period. Further, their study, like ours, integrates multiple models into a single theoretical and statistical framework.

Conceptual implications and new theoretical predictions. Given the importance of studying the consistency of patterns of associations over time (Fraley et al., 2013), we sought to test the consistency of support for predictions from each the ASC models over the critical early to middle adolescent period. The I/E model is the best-known system of knowledge that captures both the social and the dimensional comparison processes that give rise to the ASC. In contrast, the REM model represents the theoretical implications of self-concept for the critical outcome of achievement. Intriguing paradoxical hypotheses arise from the integration of these theoretical models: In the I/E model, MSC is positively predicted by math achievement but negatively predicted by verbal achievement (the dimensional comparison process). This suggests, perhaps, that verbal self-concept might also have a negative effect on subsequent math achievement (in contrast to the positive effect of MSC; Parker, Marsh, Morin, Seaton, & Van Zanden, 2015). Although this untested hypothesis appears counterintuitive, it follows directly from the theoretical integration of the I/E and the REM.

Next, we consider the place of the BFLPE within this integrated system. Unlike the REM and, perhaps, the I/E model, the BFLPE is the result of a particular event at a given point in time—namely, school selection—and thus can be considered as distinct from the integrated I/E and REM's operation within high school. Thus, unlike the REM and the I/E, the BFLPE can be seen as a response to an age-graded developmental task. Specifically, at age 10, children in Germany are sorted into different academic tracks and into schools of different average ability. This sorting of children in relation to prior achievement thus determines a child's relative position within their peer environment and subsequently, the influence of the school context on their self-concept (i.e., the BFLPE). The question then is what role this early developmental experience plays in the REM I/E system. Several possibilities are suggested by Fraley et al. (2013). First, the effect could be initially impactful, before trailing off over time. This might be the case if

initial position within the school becomes less important for self-concept over time. In this case, we would expect the effect of school average achievement on ASC to decline over time. Should this be so, the need to address the BFLPE, to arrest poor ASC and its effect on performance, might take on less practical and research significance.

Alternatively, the BFLPE could have an enduring or even increasing effect, such that school placement might assume even more prominence, due to its enduring effect on ASC. Likewise, from the perspective of the REM model, school placement might have an increasing effect, due to its influence on subsequent achievement. Furthermore, and given the I/E model, school-average achievement in one subject might have contrasting effects on ASC and achievement in other subjects.

Taken together then, this research considers the legacy of school selection (as per the BFLPE) on the system of ASC given by the integration of the I/E and REM effects. These new research issues and other benefits come from the heuristic integration of the different models into a unified conceptual framework of self-concept formation. (Also see the online Supplemental Materials Section 7 for other examples of new predictions based on the integrated ASC model.)

Methodological and design implications. Appropriate analysis of the integration of the ASC (I/E, REM, and BFLPE) models and developmental perspectives requires large, representative, longitudinal samples of students from many different schools. Particularly in BFLPE studies, appropriate multilevel models are required that also take into account the nesting of students within schools. Likewise, in REM studies, at least two and preferably many more than two waves of data are required, to test reciprocal effects between ASC and achievement, whereas in the I/E model it is important to contrast ASCs and achievement in at least two domains—typically, math and verbal. Statistical tests of developmental equilibrium require that the same set of variables be collected in at least three waves of data that span a critical developmental period of interest.

Due in part to methodological and design features that are idiosyncratic to each, research into each of these theoretical models has developed somewhat in isolation of the others. Methodologically, constraining paths to be equal (our test of developmental equilibrium) has statistical advantages (e.g., model identification, convergence, improved parsimony, increased statistical power, ease of interpretation) that are completely aside from the substantive meaning associated with support for developmental equilibrium. Although individual studies, and particularly meta-analyses of the three models of ASC, evaluate the consistency of effects over different age groups, this is rarely based on true longitudinal data in which the same set of variables is administered to the same individuals over an extended developmental period. Extending research into each of these ASC models, we posit new developmental perspectives on each through the integration of all three into a single study. From a developmental perspective, appropriate longitudinal data, strong theoretical models, and appropriate statistical analyses are important in testing the consistency of support for predictions over critical stages of development, such as the potentially turbulent years of early to middle adolescence considered here.

The Present Study: A Priori Research Hypotheses and Research Questions

Here we integrate and extend three major ASC models (REM, I/E, and BFLPE) to form an integrated ASC model in a longitudinal study of developmental equilibrium. Data (a representative sample of 3,370 students from 42 schools) were collected from the year before the start of secondary school (Year 4 school grades in German and math) and in each of the subsequent 5 years of compulsory secondary schooling in Germany (math school grades, standardized math achievement tests and MSCs). We seek to demonstrate that the three theories of self-concept formation, developmental equilibrium, and appropriate statistical methodology can be unified in a single model, as presented in Figure 1.¹

Hypothesis 1: I/E Model—Paths From Year 4 Variables to Variables in Years 5 Through 9

In Figure 1, paths (dashed lines) from primary-school math grades (Year 4) to MSC (Year 5) are predicted to be positive, but those leading from primary-school German grades (Year 4) to MSC (Year 5) are predicted to be negative (noting however, that we cannot test the corresponding paths to German self-concept, because we do not have measures of German self-concept). We leave as a research question whether there are direct effects of primary-school grades on subsequent MSCs (e.g., the effect of primary grades on MSC in Years 6 through 9 after controlling for the effects of math test scores, school grades in math, and MSCs in Year 5). However, in support of developmental equilibrium, we hypothesize an enduring legacy of the effects of achievement at the end of primary school on outcomes across secondary school years: that the total effects of primary-school grades are invariant over the multiple waves of data.

Hypothesis 2: REM—Paths Relating Variables in Years 5 Through 9

Consistently with REM research, we posit that paths leading from prior achievement to subsequent MSC (e.g., paths leading from test scores and school grades in Year 5 to MSC in Year 6) will be positive, as will be paths leading from prior MSC to subsequent achievement. In support of developmental equilibrium, these lag-1 paths (i.e., paths from variables in one wave to the immediately next wave in Figure 1; also see the Fraley et al., 2013) are predicted to be consistent over all five waves.

Hypothesis 3: BFLPE: School-Average Achievement Effects on MSC in Years 5 Through 9

Consistently with previous BFLPE theory and research, we predict that school-average achievement, operationalized in a variety of ways, will have a negative predictive effect on ASC measured in each of the first 5 years of secondary school. However, several features of our study contribute to the unique per-

¹ In presentation of the predictions and subsequent path analytic results, we use the term *effect* in its conventional statistical sense and standard path analytic terminology, as representing a relation that is not necessarily causal (see subsequent discussion of limitations in the Discussion section).

spective on this issue that extends previous research. Most BFLPE studies (Marsh, Seaton, et al., 2008; Marsh & Seaton, 2015) evaluate de facto selection (e.g., naturally occurring differences in school-average achievement on the basis of geography or post-code) on the basis of the school-average test scores at some point in secondary school, and relate this contextual variable to ASCs collected in the same school year.

As such existing BFLPE research has several potential limitations that we were able to address in this research. Thus, in our integrated ASC model (Table 1 and Figure 1), school-average math achievement is represented as a latent variable defined by three distinct contextual variables: the traditional BFLPE measure of school-average achievement based on test scores in the first year of secondary school (Year 5, based on tracked schools); school track (high, medium, or low), based on student accomplishments prior to the start of secondary school; and school-average math grades, based on the last year of primary school. However, due to

the methodological and theoretical differences associated with each of these contextual measures, we also tested each of them separately.

We hypothesize an enduring and important legacy of the negative effects of school-average achievement at the end of primary school on MSC outcomes across secondary school years. More specifically, the developmental equilibrium proposal leads to the prediction that the total effects of school-average achievement at the start of secondary school are consistent in size across the ensuing years of secondary education. However, we also note that there is an apparent clash between this prediction and previous research showing that the BFLPE increases in size the longer students are in the same school (e.g., the BFLPE should be most negative in Year 9). Hence we leave the juxtaposition of these two contrasting predictions as a research question, but note that both are consistent with the hypothesis that the BFLPE has an enduring, negative legacy.

Table 1
Integrated ASC Model: Path Coefficients Leading From Predictor Variables in Years 4 Through 8 to Math Outcomes (School Grades, Self-Concept, Test Scores) in Years 5 Through 9

Prediction of math self-concept			Prediction of math grade			Prediction of math test		
Predictor	Estimate	SE	Predictor	Estimate	SE	Predictor	Estimate	SE
MSelf5			MGrd5			MTest5		
MGrd4	.49	.04	MGrd4	.42	.03	MGrd4	.59	.03
GGrd4	-.26	.04	GGrd4	.01	.03	GGrd4	.09	.03
MSelf6			MGrd6			MTest6		
MSelf5	.45	.01	MGrd5	.47	.02	MTest5	.48	.02
MTest5	.11	.02	MTest5	.17	.02	MSelf5	.04	.01
MGrd5	.14	.02	MSelf5	.07	.01	MGrd5	.04	.01
MGrd4	.18	.04	MGrd4	.13	.03	MGrd4	.25	.03
GGrd4	-.20	.04	GGrd4	-.07	.02	GGrd4	.13	.03
MSelf7			MGrd7			MTest7		
MSelf6	.45	.01	MGrd6	.47	.02	MTest6	.48	.02
MTest6	.11	.02	MTest6	.17	.02	MSelf6	.04	.01
MGrd6	.14	.02	MSelf6	.07	.01	MGrd6	.04	.01
MGrd4	.03	.04	MGrd4	.03	.03	MGrd4	.13	.03
GGrd4	-.12	.03	GGrd4	-.02	.03	GGrd4	.06	.03
MSelf8			MGrd8			MTest8		
MSelf7	.45	.01	MGrd7	.47	.02	MTest7	.48	.02
MTest7	.11	.02	MTest7	.17	.02	MSelf7	.04	.01
MGrd7	.14	.02	MSelf7	.07	.01	MGrd7	.04	.01
MGrd4	.03	.04	MGrd4	-.04	.03	MGrd4	.11	.02
GGrd4	-.08	.04	GGrd4	-.06	.03	GGrd4	.09	.02
MSelf9			MGrd9			MTest9		
MSelf8	.45	.01	MGrd8	.47	.02	MTest8	.48	.02
MTest8	.11	.02	MTest8	.17	.02	MSelf8	.04	.01
MGrd8	.14	.02	MSelf8	.07	.01	MGrd8	.04	.01
MGrd4	.03	.04	MGrd4	-.04	.03	MGrd4	.08	.02
GGrd4	-.03	.04	GGrd4	-.02	.03	GGrd4	.05	.02
BFLPE								
Schl-Ach5	-.25	.08						
Schl-Ach	.03	.12						
Schl-Ach	.03	.09						
Schl-Ach	-.01	.07						
Schl-Ach	-.17	.07						

Note. Coefficients more than 1.96 times its standard error (values in parentheses) are statistically significant at a nominal $p < .05$. Consistently with support for developmental equilibrium, lagged paths are constrained to be invariant across the Years 5 through 9. For example, the lag-1 path $MTest_{Year5} \rightarrow MSelf_{Year6} = .447$ is equal to the corresponding lag-1 path $Mtest_{Year6} \rightarrow MSelf_{Year7} = .447$. The results from this model are used to test predictions in relation to developmental equilibrium (DevEq), the internal/external frame of reference model (I/E), the reciprocal effects model (REM), and the big-fish-little-pond effect (BFLPE). Coefficients in bold are statistically significant at $p < .05$. ASC = academic self-concept; MGrd = Math School grades (Years 4 through 9); MSelf = Math self-concept (Years 5 through 9); MTest = math test scores (Years 5 through 9); GGrd = German (Deutsch) school grades (Year 4 only); pred = predictor variables; I/E = internal/external; DevEq = developmental equilibrium.

A unique aspect of our study, arising from the integration of the BFLPE and I/E models, is the juxtaposition of effects of school-average math and verbal achievement, based on primary school grades, on subsequent MSC during the next 5 years of high school. Because both math and German contribute substantially to tracking decisions, it might be expected that the effects of school-average measures would have similar results. However, the I/E model suggests that at the individual student level, the effects of math and German grades would be in opposite directions. Extrapolating from this I/E-logic, it may be that school-average German achievement has a positive effect on MSC (in contrast to the negative effects of school-average math achievement on MSC). Hence, we leave as a research question whether the effects of school-average math grades in Year 4 on MSC differ substantially from those of school-average German in Year 4 on MSC.

Hypothesis 4: Developmental Equilibrium: Consistency of Paths Over Time (Years 5 Through 9)

Consistently with predictions based on developmental equilibrium, the critical features of our a priori model are the stability- and cross-paths in Figure 1.

1a. Autocorrelation paths. All autocorrelation (test–retest, horizontal) paths relating all Years 5 through 9 variables in each wave to the same variable in the subsequent wave (see Figure 1) are expected to be invariant across waves (e.g., all lag-1 stability paths for MSC measured in one wave to MSC measured in the next wave are constrained to be the same across all five waves).

1b. Cross-paths. Cross-paths relating all variables in Years 5 through 9 in each wave to each of the different variables in subsequent waves (see Figure 1) are also expected to be invariant across waves (e.g., lag-1 paths from test scores in Year 5 to MSC in Year 6 are the same as the lag-1 path from test scores in Year 8 to MSC in Year 9). These cross-paths in our integrated ASC model are of particular importance in testing predictions for the three theoretical ASC models that are based on the cross-paths. In particular, tests of the invariance of cross-paths provide a very strong test of developmental equilibrium and the consistency over time of predictions based on these theoretical models.

Method

Sample

The data in our study were based on the Project for the Analysis of Learning and Achievement in Mathematics (PALMA; Frenzel, Pekrun, Dicke, & Goetz, 2012; Marsh, Pekrun, Lichtenfeld, et al., 2016; Marsh, Pekrun, Parker, et al., 2016; Murayama et al., 2013; Murayama et al., 2016; Pekrun et al., 2007; Pekrun, Lichtenfeld, Marsh, Murayama, & Goetz, 2017), a large-scale longitudinal study investigating the development of math achievement and its determinants during secondary school in Germany. The study was conducted in the German federal state of Bavaria and consisted of five measurement waves spanning Years 5 to 9 in secondary school (the last 5 years of compulsory education), as well as school grades from the last year of primary school (Year 4). A unique aspect of our study that derives from the nature of the German school system is that the primary schools considered here, in contrast to the secondary schools, were not tracked, and thus were relatively heter-

ogeneous in relation to achievement in Year 4 (prior to the start of secondary school). Primarily on the basis of primary school performance, starting in Year 5, students are tracked into three school types that are relatively homogeneous in relation to achievement: high-achievement (*Gymnasium*), middle-achievement (*Realschule*), or low-achievement (*Hauptschule*) school tracks.

Excluding a small number of students for whom tracking was not introduced until Year 7 rather than Year 5, students ($N = 3,450$; 50% girls; M age = 11.7 at Year 5, $SD = 0.7$), the sampling design resulted in a representative sample of students in Bavaria, in terms of student characteristics (e.g., gender, urban vs. rural, socioeconomic status; see Pekrun et al., 2007). At the first assessment (Grade 5), the sample comprised 2001 students. In each subsequent year, the study not only tracked the children who had participated in previous assessments, but also included those children who had not yet participated in the study but had become children of PALMA classrooms at the time of the assessment (see Pekrun et al., 2007), resulting in sample sizes of 1,992 (Year 6), 2,327 (Year 7), 2,342 (Year 8), 8; and 2,461 (Year 9) 2,461. Due in part to this sampling design, a substantial portion of the sample had missing data for at least one of the measurement waves. Across the five waves, 38% participated in all five measurement waves (i.e., Grades 5 through 9), and 9%, 19%, 15%, and 19% took part in four, three, two, or one of the assessments, respectively.

Students answered the questionnaire toward the end of each successive school year. All instruments were administered in the students' classrooms by trained external test administrators. Participation in the study was voluntary and parental consent was obtained for all students. Agreement was high (100% for schools and over 90% for students at each data wave), and the final sample closely represented the intended sample and population more generally (Pekrun et al., 2007). Surveys were identified by an anonymous code number to ensure participant confidentiality.

Measures

MSC was measured in each of the five secondary school years (Years 5 through 9) with the same set of six items, using a five-point Likert scale, ranging from *not true*, *hardly true*, *somewhat true*, *largely true*, to *absolutely true*. Across the five waves, the alpha estimates of reliability were consistently high (Year 5 $\alpha = .88$; Year 6 $\alpha = .90$; Year 7 $\alpha = .89$; Year 8 $\alpha = .91$; Year 9 $\alpha = .92$). The items used to measure MSC were as follows: "In math, I am a talented student"; "It is easy for me to understand things in math"; "I can solve math problems well"; "It is easy to me to write tests/exams in math"; "It is easy to me to learn something in math"; "If the math teacher asks a question, I usually know the right answer".

Students' achievement was measured with school grades (math in Years 4 through 9; German in Year 4) and math standardized achievement test scores (Years 5 through 9). School grades were end-of-the-year final grades obtained from school documents. The standardized PALMA Math Achievement Test (Murayama et al., 2013; Pekrun et al., 2007) was based on multiple-choice and open-ended items to measure students' modeling and algorithmic competencies in arithmetic, algebra, and geometry. The test was constructed using multimatrix sampling with a balanced incomplete block design, such that the number of items increased with

each wave, varying between 60 and 90 items across the five waves, with anchor items to allow for linking the two test forms and the five measurement points. The obtained achievement scores were scaled using one-parameter logistic item response theory, confirming the unidimensionality and longitudinal invariance of the test scales (Murayama et al., 2013).

Statistical Analyses

All analyses were done with Mplus 7.3 (Muthén & Muthén, 2008–2015). We used the robust maximum likelihood estimator, which is robust against any violations of normality assumptions. All analyses were based on multilevel models (type = two level in Mplus) using manifest variables (Lüdtke, Marsh, et al., 2008; Marsh, Lüdtke, et al., 2009). Specifically, students were nested within schools, resulting in the nonindependence of observations. Ten imputed data sets were created using the default model in Mplus (see earlier discussion of the sampling design), which included all variables used in the analyses, including school-average measures of math achievement and also student background variables (i.e., socioeconomic status, gender, IQ). The final parameter estimates and fit statistics were obtained through the aggregation procedure implemented in Mplus, following Rubin's (1987) rules.

To facilitate interpretation of parameter estimates and tests of developmental equilibrium, we standardized all measures. As the standardized math test varied from year to year, test scores were standardized separately for each year. For school grades and self-concept responses that varied along a common metric, all measures were standardized in relation to values at Wave 1 (Year 5; see the online Supplemental Materials, Section 2), resulting in a standard effect size metric in relation to standard-deviation units. Particularly in longitudinal cross-lagged studies covering such a substantial period of time with many waves of data, it is important to distinguish between direct effects (the path coefficients in traditional path models) and total effects (the sum of these direct effects and the indirect effects that are mediated through intervening variables).

Developmental Equilibrium: Rationale for the Final Integrated ASC Model

In tests of developmental equilibrium, we conducted formal tests of the invariance of paths leading from one wave to the next, across all waves. To conserve space and maintain a focus on substantive issues, the detailed models in support of developmental equilibrium and related statistical issues are presented in the online Supplemental Materials, Sections 1 and 3, and summarized here briefly.

The pattern of path coefficients is determined in part by the number of lags included (see Figure 1). Thus, lag-1 paths are from a variable in one wave to a variable in the next wave, whereas lag-4 paths are from a variable in Wave 1 (Year 5) to a variable in Wave 4 (Year 9). The rationale for including paths greater than lag-1 was based on a mix of theoretical and empirical results. A priori, there is no reason why a model with only lag-1 paths should be the best model (see related discussion by Fraley et al., 2013, who also propose that models with paths greater than lag-1 should routinely be considered). If additional paths are needed to achieve

a good fit, then constraining them to be zero is likely to bias the results and the interpretation of the lag-1 paths. Hence, the theoretically more conservative approach is to include additional paths unless there is clear evidence that they are not needed.

Our final model (see Figure 1) included the following paths: from Year 4 (the last year of primary school) to all subsequent variables in the next 5 years (the first 5 years of secondary school); autocorrelation (horizontal) test-retest paths from variables (Years 5 through 9) in one wave to variables in subsequent waves (lags 1 through 4); cross-paths from measures (Years 5 through 9) of one construct to a different construct in the next two waves (lag-1 and lag-2 paths). In preliminary analyses (see Models 1 through 4 in the online Supplemental Materials, Section 3), we explored how many lags were needed to fit the data. Models with only lag-1 paths provided a poor fit to the data. Inclusion of paths from the two primary school (Year 4) variables to all variables in Years 5 through 9 (rather than only lag-1 paths to just the Year 5 variables) improved the fit. However, Model 4 (with lags 1 through 4 autocorrelation paths but only lag-1 and lag-2 cross paths) provided an excellent fit to the data (see the online Supplemental Materials, Section 3); this is consistent with our supposition that more than lag-1 paths are needed. This model with multiple lags is also conservative, providing stronger controls for preexisting differences, particularly compared with the typical approach used in developmental studies, which are based on only two waves of data, and studies with more than two waves of data that ignore paths other than lag-1 effects.

In Model 5 (see the online Supplemental Materials, Section 3), we added invariance constraints to Model 4 to test the a priori assumption of developmental equilibrium (that the paths are consistent over waves). In Model 5, there was strong support for complete invariance of autocorrelation- and cross-paths, across all waves. This included the invariance of cross-paths (e.g., MSC wave_{*i*} → Test Scores wave_{*i+1*} = MSC wave_{*i+1*} → Test scores_{*i+2*}), which are central to tests of predictions from our three theoretical models of ASC formation. However, autocorrelation paths for Lag 1 through 4 paths were also shown to be invariant. For example, not only were lag-1 paths invariant (e.g., MSC wave_{*i*} → MSC wave_{*i+1*} = MSC wave_{*i+1*} → MSC wave_{*i+2*}), but also lag-2 paths (e.g., MSC wave_{*i*} → MSC wave_{*i+2*} = MSC wave_{*i+1*} → MSC wave_{*i+3*}), lag-3 paths and lag-4 paths were also invariant. The fit of Model 5 was excellent, and differed little from that of Model 4 with no invariance constraints; this provides support for the more parsimonious model. In the final model in this sequence (i.e. the integrated ASC model; see the online Supplemental Materials, Section 3), we added three school-level variables (school-average math grades from Year 4, school track, and school-average test scores) that defined a latent variable used to infer school-average achievement. This model provided an excellent fit to the data and provided the basis for all subsequent analyses. This latent factor of school-average achievement was well-defined (e.g., standardized factor loadings for the three indicators varied from .94 to .98) and inclusion of the factor at the school level actually resulted in a marginally better fit to the data, relative to the model with no school-level factors (see the online Supplemental Materials, Section 3). Finally, we repeated all the analyses leading up to the final model, including the latent school-average achievement factor to the final model (see the online Supplemental Materials, Section 3), demonstrating that this did not alter conclusions about the structure of the model at the student level.

This final multilevel model (see Figure 1 and Table 1; also see the online Supplemental Materials, Section 3, for further discussion) is referred to as the integrated ASC model because parameter estimates from this one statistical model provide tests of a priori predictions in relation to the internal/external frame of reference model (Hypothesis 1), the reciprocal effects model (Hypothesis 2), and the big-fish-little-pond effect (Hypothesis 3). Support for the invariance of path coefficients over Years 5 through 9 (the first 5 years of secondary school), demonstrated through tests of developmental equilibrium, provides an important developmental perspective on each of the ASC models considered here. Thus, the invariance of parameter estimates over time provides support for developmental equilibrium in respect of each of the ASC models (Hypotheses 1 through 3), as well developmental equilibrium more generally.

Results

The latent correlation matrix among student-level variables (see the online Supplemental Materials, Section 2) provides an advanced organizer in relation to subsequent analyses. Math test scores in secondary school (Years 5 through 9) were substantially and consistently correlated across the 5 years for both primary school (Year 4) math grades ($r_s = .65$ to $.70$) and German grades ($r_s = .48$ to $.57$). Interestingly, these correlations of primary school grades with secondary school test scores were substantially higher than corresponding correlations between test scores and school grades in secondary school (.26 to .43). This is consistent with earlier discussion, and suggests that school grades in untracked primary schools are more like test scores, in that they reflect a more common underlying metric continuum than do grades in Years 5 through 9 in the highly tracked secondary schools. These results have potentially important implications for understanding the grading on a curve phenomenon (see Supplemental Materials, Section 2, for further discussion), as well as our subsequent use of school-average achievement based on these Year 4 math grades. Finally, consistent with much previous research, correlations between MSC and school grades ($r_s = .42$ to $.62$) are higher than the corresponding correlations between MSCs and test scores ($r_s = .30$ to $.34$).

Frame of Reference Effects: The I/E Model (Hypothesis 1)

In tests of the I/E model, our focus is on the dashed paths in Figure 1, reflecting the effects of primary school grades (Year 4) on subsequent measures of MSC collected during secondary school (Years 5 through 9). Tests of this model are based on our integrated ASC model (see Table 2). Consistent with the I/E model and Hypothesis 1, the path from Year 4 math grades to MSC in Year 5 is significantly positive (.49, Table 2), whereas the path from Year 4 German grades to Year 5 MSC is significantly negative (-.26).

As noted previously, in evaluating the consistency of effects of Year 4 grades over MSC in Years 5 through 9 we focus on total effects (but present both total and direct effects in Table 2). The total effects of Year 4 Math grades on MSCs in Years 5 through 9 were consistently positive (paths = +.48 to +.54) and were not significantly different from each other (Wald test [$df = 4$] =

Table 2
Tests of the Internal/External Frame of Reference Model

Year	Predictor		Outcome prediction					
			Math self		Math grade		Math test	
			Estimate	SE	Estimate	SE	Estimate	SE
5	Total	MGrd4	.49	.04	.42	.03	.59	.03
		GGrd4	-.26	.04	.01	.03	.09	.03
	Direct	MGrd4	.49	.04	.42	.03	.59	.03
		GGrd4	-.26	.04	.01	.03	.09	.03
6	Total	MGrd4	.54	.05	.46	.03	.57	.03
		GGrd4	-.29	.05	-.07	.03	.16	.03
	Direct	MGrd4	.18	.04	.13	.03	.25	.03
		GGrd4	-.20	.04	-.07	.02	.13	.03
7	Total	MGrd4	.48	.04	.38	.04	.56	.04
		GGrd4	-.28	.04	-.04	.04	.15	.04
	Direct	MGrd4	.03	.04	.03	.03	.13	.03
		GGrd4	-.12	.03	-.02	.03	.06	.03
8	Total	MGrd4	.48	.05	.34	.03	.58	.03
		GGrd4	-.27	.04	-.09	.03	.19	.03
	Direct	MGrd4	.03	.04	-.04	.03	.11	.02
		GGrd4	-.08	.04	-.06	.03	.09	.02
9	Total	MGrd4	.49	.04	.34	.03	.58	.03
		GGrd4	-.25	.05	-.06	.03	.18	.03
	Direct	MGrd4	.03	.04	-.04	.03	.08	.02
		GGrd4	-.03	.04	-.02	.03	.05	.02

Note. Mgrd4 = Math school grades in Year 4. Ggrd4 = German school grades in Year 4. Effects (standard errors in parentheses) are for German and math school grades from primary school (Year 4) on secondary school outcomes in Years 5 through 9 (math self-concept, math school grades, and math test scores). Total effects include both direct and indirect effects. Thus, for example, the total effects of Year 4 math grades on Math grades in Years 5 through 9 are consistently substantial (.459 to .420), but direct effects are only significant for Years 5 and 6; most of the effects of Year 4 Math school grades on Math grades in Years 7 through 9 are mediated through math grades in Years 5 and 6. Coefficients in bold are statistically significant at $p < .05$. Also see Figure 1 and Table 1.

8.195, $p = .085$). In contrast, the total effects of Year 4 German grades on MSCs (paths = $-.25$ to $-.29$) in Years 5 through 9 were consistently negative, but again were not significantly different from each other (Wald test [$df = 4$] = 1.179, $p = .703$). In summary, there is strong support for the I/E model and for the developmental equilibrium hypothesis, on the basis of the consistency of the effects of Year 4 grades across MSC in Years 5 through 9.

It is also informative to evaluate the direct effects of Year 4 grades on Years 5 through 9 outcomes, controlling for intervening variables (noting that both the direct and total effects are based on Model 7). The question then becomes what are the direct effects of Year 4 school grades on, for example, Year 9 MSC after controlling for the indirect effects that are mediated through MSC, math school grades, and math test scores from Years 5 through 8 (the direct paths in Table 2 are the same as those presented in Table 1, but are repeated to facilitate the comparison of direct and total effects in Table 2). Beyond Year 5, the direct effects of Year 4 math grades continue to be significantly positive for Year 6 MSC (.18), in addition to the positive effects on Year 5 MSC (path = .49), whereas paths to MSC in Years 7 through 9 are nonsignificant. Of course, in subsequent Years 6 through 9, there are still substantial total effects of Year 4 math grades, but these are mediated through outcomes in intervening years.

The pattern of results based on Year 4 German grades is different. In addition to effects on MSC at Year 5, there are new, statistically significant negative direct effects on MSCs in Years 5 through 8 (-.20, -.12, and -.08, respectively). It is only in Year 9 that the negative effects of Year 4 German grades are no longer statistically significant. These new effects of Year 4 German grades in Years 7 and 8 are likely to be due to the fact that intervening variables during secondary school years did not include measures of German achievement, which would otherwise mediate the effects of Year 4 German grades, as was the case with Year 4 math grades.

REM: Temporal Ordering of School Grades, Test Scores, and MSC (Hypothesis 2)

In evaluating the temporal ordering of ASC and achievement, we focus on lag-1 cross-paths relating MSC and math achievement (see Figure 1). The results are consistent with the REM (Hypothesis 2): For our integrated ASC model (see Table 3; also see Figure 1 and Table 1), paths are positive leading from prior MSC to subsequent achievement (school grades, .07, SE = .01; and test scores, .04, SE = .01) and from prior achievement to subsequent MSC (school grades, .14, SE = .02; and test scores, .11, SE = .02). The reciprocal effects associated with school grades were significantly higher than were those for test scores (Wald = 9.761, df = 2, p = .008). This also is consistent with previous research,

although these differences were surprisingly small (see Table 3). In additional models (see Table 3) we also tested the effects of grades (without test scores) and test scores (without grades). Not surprisingly, when both school grades and test scores were considered separately, the sizes of the reciprocal effects between achievement and MSC were somewhat higher. However, the pattern of results was similar, in that all reciprocal effects were significantly positive.

BFLPE: Negative Effects of School-Average Achievement (Hypothesis 3)

In support of this hypothesis and extensions of the BFLPE over the five secondary school Years 5 through 9 (see Figure 1), we begin with the integrated ASC model (see Table 4) in which school-average achievement is based on a latent factor defined by school-average math Year 4 grades, school-average test scores, and school track. This latent factor was well-defined, as correlations among the three indicators varied from .89 to .94, and the standardized factor loadings varied from .94 to .98. The direct effects were significant in Year 5, but were not significant in Years 6 through 8, indicating that the initial negative effects in Year 5 neither increased nor decreased during these years. However, there was a significantly negative direct effect in Year 9, indicating that the BFLPE is significantly more negative in Year 9, even after controlling for the negative effects in Years 5 through 8. This indicates that following the significant negative effect in Year 5, there is no significant change in the size of the BFLPE over the period of Years 6 through 8, but there is a new, significantly negative BFLPE at Year 9 beyond what can be explained in terms of the effects in Years 5 through 8.

This pattern of results for the direct effects is consistent with the finding that the total effects were statistically significant in all Years 5 through 9 (-.13 to -.26). The test of the equality of the total effects was rejected (Wald = 21.73, df = 4, p < .001), but orthogonal polynomial contrasts were nonsignificant for both linear and quadratic effects. However, comparisons of the total effect for each year against the mean effects for the other 4 years revealed only one statistically significant difference: the negative BFLPE was significantly smaller at Year 8 (-.13, p = .008; see Table 3). Interestingly, although the size of the BFLPE was most negative in Year 9 (-.31), this effect was not significantly different from the average total effects in Years 5 through 8 (p = .132). In summary, on the basis of the total effects, the direction of the BFLPE was reasonably consistent over the 5 years. Although the most negative effect was in the final year of mandatory school, the nonsignificant linear trend suggests that the effects were not systematically increasing or decreasing over this critical developmental period. Hence, there is clear support for Hypothesis 3: that there is an enduring negative legacy of the BFLPE.

Effects of different indicators of school-average math achievement and school track. In Models A1 through A3 (see Table 3), we tested the BFLPE separately for: school-average achievement based on Year 4 math grades (Model A1), school track (Model A2), and math test scores (Model A3). Not surprisingly, given the very high correlations among these different indicators, the pattern of results based on each is similar to that which is based on the latent factor already discussed. However, this consistency across the three indicators is important, as two of

Table 3
Tests of the Reciprocal Effects Model Based on Integrated Model: Lag-1 Paths From Each Predictor Variable in One Year to Each Outcome in the Next Year for Math Self-Concepts, Grades, and Test Scores in Secondary School Years 5 Through 9

Predictor	Outcome prediction					
	Math self		Math grade		Math test	
	Estimate	SE	Estimate	SE	Estimate	SE
Integrated ASC model (grades and test scores)						
Math self-concept	.45	.01	.07	.01	.04	.01
Math grades	.14	.02	.47	.02	.04	.01
Math test	.11	.02	.17	.02	.48	.02
Predictor grades only						
Math self-concept	.48	.01	.10	.01		
Math grades	.16	.01	.51	.02		
Predictor test only						
Math self-concept					.05	.01
Math test	.13	.02			.62	.02

Note. Estimates (standard errors in parentheses) for the integrated academic self-concept (ASC) model. For the reciprocal effects model, critical paths in relation to a priori redictions are the cross-paths leading from each predictor variable in one wave to a different outcome variable in the next wave (e.g., Mtest_{Year5} → MSC_{Year6} where M = math, SC = self-concept). Horizontal paths are paths from each variable to the same variable in a subsequent wave (e.g., Mtest_{Year5} → MSC_{Year6}). Consistently with support for developmental equilibrium in the integrated ASC model, lag-1 paths are constrained to be invariant across Years 5 through 9 (e.g., Mtest_{Year5} → MSC_{Year6} = Mtest_{Year6} → MSC_{Year7}). In variations to the integrated ASC model specific to the reciprocal effects model, we tested separately for math grades and standardized tests. Coefficients in bold are statistically significant at p < .05.

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Table 4
Big-Fish-Little Pond Effect (BFLPE): Effects of School-Average Variables on Math Self-Concept in Years 5 Through 9

Dependent	Common model		Model A1		Model A2		Model A3		Model B1	
	Latent factor		School average		School		School average		School average	
	(A1 + A2 + A3)		Math grades		Track		Math test		German grades	
	Estimate	SE	Estimate	SE	Estimate	SE	Estimate	SE	Estimate	SE
MSC-Year 5 Total	-.25	.08	-.18	.06	-.16	.05	-.23	.07	-.20	.06
Direct	-.25	.08	-.18	.06	-.16	.05	-.23	.07	-.20	.06
MSC-Year 6 Total	-.26	.08	-.19	.06	-.17	.05	-.23	.07	-.20	.06
Direct	.03	.12	.02	.08	.01	.07	.04	.10	.04	.10
MSC-Year 7 Total	-.18	.08	-.13	.06	-.10	.05	-.18	.07	-.12	.06
Direct	.03	.09	.03	.06	.04	.05	-.02	.08	.05	.07
MSC-Year 8 Total	-.13	.06	-.10	.05	-.07	.04	-.14	.06	-.10	.05
Direct	-.01	.07	.00	.05	.00	.04	.00	.05	-.02	.05
MSC-Year 9 Total	-.31	.09	-.22	.07	-.19	.06	-.28	.08	-.22	.07
Direct	-.17	.07	-.11	.05	-.11	.04	-.14	.06	-.11	.05
Mean total effect	-.23		-.16		-.14		-.21		-.17	

Note. MSC = Math self-concept (Years 5 -9). Estimates (standard errors in parentheses) for the direct paths from school-average achievement come from the integrated academic self-concept (ASC) model and the corresponding total effects. In the integrated ASC model, school-average achievement represented by one latent variable (based on school-average values for math grades, school track, and test scores). In additional models specific to tests of the BFLPE, separate analyses were done for each of these three measures of school-average math achievement: School average math grades (Model A1, based on Year 4, the last year of primary school); school track (Model A2: high, medium, low track determined prior to the start of secondary school); and school-average test scores (Model A3, based on test scores from Year 5, the first year of secondary school). Model B1 tests the effects of school-average German grades. Coefficients in bold are statistically significant at $p < .05$. MSC = math self-concept; Year 5 through Year 9 = first five years in secondary school.

the indicators are based on primary school performance prior to the start of secondary school. In contrast to these true pretreatment indicators, most previous BFLPE studies are based on posttreatment measures of achievement—as with our school-average measure of math test scores.

Next, we juxtaposed the effects of school-average achievement based on math grades (Model A1) with those based on German grades (Model B1). The critical finding is that the negative BFLPEs for school-average German grades are nearly the same as those based on math grades. Again, this is not surprising, in that the correlation between these two school-average measures ($r = .983$) is very high, such that contextual effects associated with math and verbal performances could not be distinguished.

Discussion

The present investigation is a large, longitudinal panel study covering the period from the end of primary school through the first 5 years of secondary school (the end of compulsory schooling in Germany). The study uniquely integrates three of the main theoretical models in ASC research (REM, I/E, BFLPE) and offers new developmental perspectives on each. It does so, in part, by capitalizing on a characteristic feature of the German school system, in which students attend nonselective primary schools up to Year 4, but subsequently attend highly achievement-tracked secondary schools. A critical feature of this study is the longitudinal design, covering the last year of primary school and the first 5 years of secondary school, which provides a unique developmental perspective on the consistency and robustness of effects over this critical early to middle adolescent period. In particular, there was strong support for a very demanding test of developmental equilibrium, demonstrating that support for each of the three theoretical

ASC models was highly robust in the developmental period. Importantly, this support for developmental equilibrium is based on an overarching conceptual and methodological framework in which all the predictions are tested in relation to parameters from a single statistical model. The integration is also heuristic in providing new research questions, some of which were tested here, while others remain directions for further research. We now provide a summary of critical new substantive findings, emphasizing the significance of this integration of the three theoretical models and developmental equilibrium into a common conceptual and methodological framework.

Integration of the I/E Model, REM, and Developmental Equilibrium

Tests of the I/E model are traditionally based on a single wave of data and focus on paths leading from achievement to ASC, thus confounding the temporal ordering of ASC and achievement, which is a salient feature of the REM. The integration of the I/E model with the REM and developmental equilibrium resolved this issue in demonstrating the effects of math and German school grades from the end of primary school on math constructs over the next 5 years of secondary school. The effects of math grades on subsequent MSCs were substantial and positive, whereas the effects of German grades on subsequent MSCs were substantial and negative. This is consistent with the I/E model. In keeping with our developmental equilibrium perspective, the total effects of Year 4 grades on MSCs were remarkably consistent over the next 5 years of secondary school. The consistency of these effects over such an extended developmental period is apparently unique in I/E studies; it suggests that the I/E effects are robust and that this pattern of relations has achieved developmental equilibrium over this period.

The temporal ordering of academic achievement and ASC has important theoretical and policy-practice implications, but most REM studies are based on only two waves of data, neither of which provides tests of developmental equilibrium or of generalizability over critical developmental periods. Our investigation is one of the few to have considered as many as six waves of data—including school grades from before the start of high school. In support of developmental equilibrium, the effect of MSC in one school year to achievement (school grades and test scores) in the next year is similar across all five waves, as is the effect of achievement on MSCs in the following wave.

This pattern of results in support of the REM was consistent over test scores and school grades considered separately or in combination. Although the reciprocal effects associated with school grades were significantly stronger than were those for test scores, it is surprising that these differences were not even larger. Indeed, inspection of the correlations (see the online Supplemental Materials, Section 2) shows that for all Years 5 through 9, MSCs are substantially more correlated with school grades (.42 to .62) than with test scores (.30 to .34). The apparent explanation is that controlling for school grades at Year 4 reduced the effects of subsequent school grades more than the effects of subsequent test scores. Hence, in unreported, supplemental analyses that did not control for Year 4 grades, the reciprocal effects associated with both grades and test scores increased substantially, but the increases for reciprocal effects associated with grades were substantially larger than were those associated with test scores.

A limitation of the present investigation, in relation to this integration of the I/E, REM, and developmental equilibrium, is that German achievement and self-concepts were not collected in Years 5 through 9. Thus, the positive effects of Year 4 math grades on MSCs in Years 5 through 9 were remarkably robust, as were the negative effects of Year 4 German grades on MSCs in Years 5 through 9. However, it was not possible to test the corresponding predictions about the effects of math and German grades on German ASCs in Years 5 through 9.

Importantly, integration of the I/E model and REM suggests a provocative new prediction that we were unable to test with the data available. The REM highlights the reciprocal effects of achievement and self-concept in matching domains, while the I/E model emphasizes effects of achievement on self-concept in contrasting domains that are negative. Integrating the two suggests that prior self-concept in one domain should have a negative effect on subsequent achievement in a contrasting domain. Although tests of this prediction must be left for future research (but see Möller, Retelsdorf, Köller, & Marsh, 2011), as it is not testable with the available data, this new prediction demonstrates the heuristic value of the integration of the different theoretical models (also see the online Supplemental Materials, Section 7).

Extending the BFLPE and Its Integration With the I/E Model, REM, and Developmental Equilibrium

Our study provides new perspectives on the potential limitations of previous BFLPE studies. These include the focus on explicit (rather than de facto) tracking, the inclusion of multiple contextual variables reflecting school-average achievement, and the juxtaposition of BFLPEs based on school-average achievement inferred

from school grades with those inferred from standardized test scores. Of particular importance, our longitudinal analyses and our focus on developmental equilibrium have demonstrated the generalizability of BFLPEs over time (BFLPEs in Years 5 through 9) and over the different contextual variables used to represent school-average achievement.

The juxtaposition between these three contextual measures of school-average achievement is important in addressing a potential confounding of the temporal ordering of ASC and achievement in most BFLPE studies. In particular, because most BFLPE studies are based on a single wave of data, school-average achievement typically is inferred from test scores collected at the same time as the ASC measures. In our investigation, this was the case with BFLPEs based on school-average test scores from Year 5 (collected after students had already started secondary school). However, we showed that these potentially confounded effects on this contextual variable were consistent with those for school-average achievement based on Year 4 school grades and school track, true pretreatment variables determined before students began their secondary schooling. This consistency was also evident in the substantial correlations among the three contextual variables (r s of .89 to .94), which allowed us to combine them into a single latent factor (i.e., the integrated ASC model). BFLPE estimates based on this latent factor were somewhat larger than were BFLPEs based on any of the measures considered separately. In summary, our study is apparently the first BFLPE study to use school grades prior to the start of secondary school to assess contextual effects, but also the first to demonstrate support for a latent school-average factor based on such apparently diverse measures of school-average achievement.

The integration of the BFLPE with developmental equilibrium also addresses a critical issue about whether the size of the BFLPE diminishes over time, remains stable, or actually becomes more negative as students progress through school. Previous research suggests that the BFLPE grows more negative over time, but this research is typically based on a limited time span or multiple, cross-sectional cohorts. In contrast, developmental equilibrium suggests that the BFLPE should remain consistent over time. Our longitudinal study, based on six waves of data, was uniquely suited to address this issue. In line with a developmental equilibrium perspective, the direction and even the sizes of the BFLPEs were reasonably consistent over Years 5 through 9. Although linear and quadratic effects were nonsignificant, a test of equality of the BFLPEs over time was rejected. In keeping with previous research suggesting that the BFLPE becomes more negative over time, the most negative BFLPE was for Year 9, the last year of compulsory education. However, the BFLPE at Year 9 was not significantly different from the average BFLPE for Years 5 through 8. Hence, there was no clear resolution as to whether the BFLPE remains stable (developmental equilibrium) or becomes increasingly negative over time (as suggested by previous research). However, it is apparent that the strength of the BFLPE does not diminish over time. In summary, the BFLPE is highly robust, with enduring negative effects across the first 5 years of secondary school.

The integration of the BFLPE model and the I/E model provides an interesting new prediction about the effects of school-average math and German school grades on MSC. At the level of the individual student in the multilevel model, support for the I/E model shows that the effects based on math and German school

grades on MSC go in opposite directions (positive for math grades and negative for German grades). Extending this I/E logic to the level of school-average achievement, there might also be counterbalancing effects of school-average math and German achievement on MSC (negative for school-average math achievement, positive for school-average German achievement). However, these two contextual variables reflect similar processes, in that math and German achievement are substantially correlated at the level of the individual student, and the track to which a student is assigned is based substantially on both math and German achievement. Thus, the school-average measures based on math and German school grades are so highly correlated ($r = .985$) that these contextual effects cannot be readily distinguished. However, as suggested by Marsh (1991), better support might be found in studies where selection into magnet schools is based on achievement in one narrowly defined academic domain (e.g., math, science, literature, sport, performing arts; see Marsh & Roche, 1996; Parker, Marsh, Lüdtke, & Trautwein, 2013; Trautwein, Gerlach, & Lüdtke, 2008) that is made highly salient to students, rather than a global measure that encompasses a diverse range of different domains. Hence, whereas here we found no support for this prediction based on the integration of the I/E and BFLPE perspectives, the rationale warrants consideration in different settings where school-average achievement in different academic domains is not so highly correlated.

Integrating ASC Models and Developmental Equilibrium into a Socio-Ecological Systems Model

The integrated ASC model can also be analyzed in terms of a broad socioecological systems model (e.g., Bronfenbrenner, 1979). This paradigm shows not only how the integration works at different levels, but how it results in new predictions at different levels of the system. At the *exosystem* level, Germany's system provides a critical test case for self-concept research, evidenced by the amount and quality of research on the topic emerging from the country (e.g., Marsh, Köller, & Baumert, 2001; also see review by Marsh & Seaton, 2015) and by the fact that BFLPEs are largest in tracking countries like Germany (Nagengast & Marsh, 2012; Salchegger, 2016). The effect of tracking on children is mediated down to individuals through school selection at the microsystem level through the BFLPE, which is intensified in countries with highly stratified school systems. At the individual level the BFLPE has its influence directly on the external comparison aspect of the I/E model. This in turn has an influence on the individual chronosphere level via the REM model, in which ASC but also achievement subsequently feedback up to microsystem, exosystem, and macrosystem (i.e., student performance and well-being ultimately are inputs into government decision processes about how to structure educational systems). Our article then explores how this developmental system as a whole operates and whether it is stable across a critical developmental period (developmental equilibrium) and thus largely immune to maturation effects.

Strengths, Limitations, Directions for Further Research, and Implications for Practice

The present investigation had a number of strengths that provide directions for future research, particularly research that seeks to

integrate predictions and developmental perspectives based on the I/E, REM, and BFLPE models. Of particular importance were: an appropriate data set with multiple waves of data (preferably three or more), multiple content domains (typically math and verbal, but possibly more than two), and representative samples of students from a large number of different schools. Although the focus of this study was on developmental perspectives in relation to the integration of three theoretical models of self-concept formation, like Fraley et al. (2013) we suggest that this synergy of strong theoretical models, suitable longitudinal data, and appropriate statistical models, provides a more general framework in which to study patterns of associations across time that address critical issues in developmental science. Of course, this framework is flexible in relation to variations that might be useful in examining issues other than those addressed here. Thus, for example, it is possible to relax the equality constraints over time, so as to introduce more complex processes (e.g., so-called sleeper effects in which the effects increase over time). More generally, these conceptual frameworks and statistical models could also be useful for studying entirely different constructs and theoretical models.

Our integrated ASC model also provides an opportunity to consider new predictions that are not evident in any of the component models considered in isolation (for further discussion, see the online Supplemental Materials, Section 7). For example, how does the internal component of the I/E model operate within this system? The REM posits that ASC has reciprocal effects with achievement in matching areas, but the internal comparison process of the I/E model suggests that the effects of ASC might be negative for achievement in nonmatching areas (e.g., effect of MSC on verbal achievement). The BFLPE posits that school-average achievement has a negative effect on ASC in the matching area, but the internal comparison process of the I/E model suggests that the effects of school-average achievement might be positive for achievement in nonmatching areas (e.g., effect of school-average verbal achievement on MSC).

In relation to potential weaknesses and directions for further research, our theoretical models posit causal effects. Although it is appropriate to make causal predictions, it is important to emphasize that our structural equation models do not conclusive fully test causality (see related discussion by Fraley et al., 2013). In the absence of random assignment with experimental manipulation, or even when there is random assignment, typically there are alternative explanations of the effects that may differ from implicit causal interpretations. Thus we use the term "effect" in its conventional statistical sense, as representing a relation that is not necessarily causal; we specifically avoid making the imputation that our effects are causal. Nevertheless, it is relevant to note that the design of our longitudinal study, with six waves of data, including a primary school measure of achievement, is stronger than most previous nonexperimental research. Particularly in relation to interpretation of the BFLPE, primary school grades and the tracking variable constitute true preintervention variables in relation to the move from untracked primary schools to highly tracked, ability-stratified high schools. This quasi-experimental design is apparently stronger than previous nonexperimental BFLPE research, and alternative designs that do not involve random assignment. Finally, we also note that the conceptual frameworks and statistical models developed here would also be useful in evaluating true intervention studies in which there

is random assignment to conditions. Thus, for example, essentially the same model would be appropriate in a hypothetical true experimental study in which students were randomly assigned to students to schools differing in levels of academic achievement.

Here we used the manifest model of the BFLPE as described by Marsh, et al. (2009), where, even for the simplest contextual models, at least 50 and preferably as many as 100 schools are recommended. Where the number of schools is relatively small (less than 50, as in our study), they found that the more parsimonious manifest models are likely to be more accurate. Nevertheless, it would be desirable for future studies to include more schools than were considered here.

Our focus was on the BFLPE and ASC over this early to middle adolescent period. However, there is also a need to extend this research to include other developmental periods (e.g., early to middle childhood, where tests of developmental equilibrium might not be supported), to expand the tests to include data beyond those from secondary school, and to consider other outcomes, including long-term effects on academic achievement, aspirations, and educational attainment. Also, it would be interesting to see the interplay of mechanisms underlying the three models of ASC over transitional periods other than the primary-to-secondary school transition considered here (e.g., secondary school to university, or from school to work). Finally, although some specific features of the German school system (particularly in relation to tracking) lent strengths to the present investigation, there is also a need to test the generalizability of the results in other countries and different school systems, and to test the extent to which results generalize across individual student and school-level characteristics within each study.

In respect of the nature of school grades, and of grading on a curve, the results provide an important demonstration that school grades can, in some circumstances, provide a common metric across different teachers and schools (see further discussion in the online Supplemental Materials, Section 1). However, our results are highly dependent on circumstances that, at least in part, are specific to the German school system. Hence, further research is needed to test the generalizability of these results and to improve the usefulness of school grades as a measure of achievement that is generalizable across teachers, school subjects, and schools.

In conclusion, in a rapidly changing world, the development of positive ASCs is important. Thus, for example, Marsh and Yeung (1997) demonstrated that ASC predicted subsequent coursework selection better than did corresponding measures of achievement, while Marsh and O'Mara (2008) showed that ASC formed in high school contributed to the prediction of long-term educational attainment 8 years later, beyond the effects of school grades, standardized achievement tests, IQ, and socioeconomic status. However, despite the tremendous growth and sophistication of ASC research in the last 30 years, major theoretical models (REM, I/E, and BFLPE) have tended to develop in isolation of each other, due in part to the methodological design complications associated with each. Here we derived an integrative set of theoretical hypotheses that spans all three models, and offered new developmental perspectives in relation to tests of developmental equilibrium.

From a developmental perspective, we have expanded the theoretical and statistical rationale for tests of developmental equilibrium, given that our longitudinal data cover the potentially turbu-

lent early to middle adolescent period. This perspective on the consistency of effects from one wave to the next is different from the typical developmental focus on the latent mean trajectories of a given construct in a period of developmental significance. Unlike many developmental studies, our focus was on developmental consistency of effects over time, rather than on developmental differences. More specifically, we found support for developmental equilibrium in terms of the invariance of effects across five waves of data, based on the assumption that the self-system had attained a developmental balance in relation to predictions from each of our theoretical models of ASC formation. Although developmental equilibrium is a powerful lens to bring a developmental perspective to each of these theoretical models, it is important to emphasize that these models do not per se depend on developmental equilibrium: the consistency of effects over a developmental period of interest. Developmental equilibrium does, however, provide an important test of whether the dynamics underlying the formation of ASC in these models have reached a state of balance over the developmental period under consideration.

In extending tests of developmental equilibrium, we offer new developmental perspectives on each of the three ASC models, including the need for longitudinal data based on more than just 2 or 3 waves, and for stronger tests of the statistical assumptions underlying the models. Indeed, we suspect that our finding of significant effects associated with lag-2, lag-3, and even lag-4 autocorrelation effects would be likely in many developmental studies if this was tested formally on the basis of a sufficient number of waves. From a statistical and design perspective, this finding is important, because failure to consider these higher order lags (because, e.g., only two waves of data were collected, or the lags were simply ignored) is likely to positively bias results, unless there is strong evidence that they are insignificant. Such results might be merely a statistical issue, in that the results of two measures of the same construct are likely to better predict a third measure than either one considered separately. However, although this is beyond the scope of the present investigation, there might also be substantively important implications of the higher order lagged effects. For example, materials taught and tested in Year 5 might lead to materials being covered in Year 7 but not tested in Year 6, so that mastery of materials in Year 5 contributes to prediction of achievement in Year 7 beyond what can be explained in terms of achievement in Year 6. More research is needed to disentangle statistical from substantive issues in the interpretation of lag-2, lag-3, and even lag-4 autocorrelation effects. Nevertheless, the default assumption should be that these higher order lagged effects do exist, in the absence of empirical tests showing that they are nonsignificant.

From the perspective of educational practice, we found that the feedback students received, in the form of teacher-assigned grades, reciprocally influenced the development of their ASCs over 5 years. However, grading that uses normative standards, such as grading on the curve, inevitably implies that some students can experience success, whereas others must fail; accordingly, the findings suggest that normative grading should be replaced by absolute or individual standards wherever possible, to promote students' ASCs. Indeed, support for the REM suggests why the job of the classroom teacher is so difficult; they need to teach academic skills, but also to reinforce positive self-beliefs and link the two in a way that is consistent with their reciprocal relations, in

which each contributes to the other. Finally, the effects of school-average achievement (defined by test scores or grades) and school track on ASC were consistently negative over the secondary school years. Accordingly, parents and teachers are well advised to consider ASC implications when reaching decisions about how best to select schools and classes for their children.

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